# Analysis of Convergence Rate for Biogeography Based Optimization 

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#### Abstract

With the dramatic improvement of computation ability in modern computers, a lot of new metaheuristic methods appear. All these meta-heuristic algorithms are originated from some mechanisms in nature, which are alike in composition and extensively adopted to solve combination optimization problems. However, meta-heuristic algorithm, such as BBO, is lack of strict theory foundation and hard to be analyzed in theory, because it comes from heuristic idea and has complicated random behavior. Therefore, in this paper, we propose some theoretical analysis to lay solid foundation for BBO application. To obtain the average first hitting time (AFHT) of BBO, firstly the relationship between AFHT and convergence rate is discussed, and then the upper bound and the lower bound of convergence rate are put forward. When BBO is adopted to solve an optimization problem with problem size n and population size n , if migration operation and Mutation-Constant or Mutation$1 / \mathrm{n}$ operation are adopted, then the lower bound of AFHT is proposed.


Keywords Meta-Heuristic, Global Optimization, NP Hard Problem.

## Introduction

Since 1960s, a lot of researchers are interested in evolutionary computation with the development of computer technology. Some famous algorithms are proposed, such as Genetic Algorithm, Evolutionary Programming, and Evolutionary Strategy. Since the computation ability of computer improves dramatically, a lot of new metaheuristic methods arise, such as Ant Colony Optimization (ACO) [1], Particle swarm Optimization (PSO) [2], Differential Evolution (DE) [3], Grey Wolf Optimization (GWO) [4], Biogeography Based Optimization (BBO) [5], Hybrid Grey Wolf Optimization (HGWO) [6] ,et al. All those algorithms are originated from some mechanisms in nature, and are similar in structure. We call such algorithms as evolutionary algorithms. In general, evolutionary algorithms are random and heuristic optimization methods, which are widely used to solve combination optimization problems. As we use evolutionary algorithms to solve optimization problems, it is easy to add some heuristic ideas. Therefore, evolutionary algorithms demonstrate excellent performance in a lot of fields such as data mining, engineering optimization an industry design.
However, evolutionary algorithm, such as BBO, is lack of strict theory foundation and hard to be analyzed in theory [7], because it comes from heuristic idea and has complicated random behavior. In this paper, we propose some theoretical analysis to lay solid foundation for BBO application. To obtain the average first hitting time (AFHT) of BBO, firstly the relationship between AFHT and convergence rate is discussed, and then the upper bound and the lower bound of convergence rate are put forward. When BBO is adopted to solve an optimization problem with problem size $n$ and population size $n$, if migration operation and Mutation-Constant or Mutation$1 / \mathrm{n}$ operation are adopted, then the lower bound of AFHT is proposed.
Organization of the rest of this paper is as follows. The basic of BBO is briefly introduced in Section 2. Analysis of Convergence rate is presented in Section 3. Finally, concluding remarks is given in Section 4.

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## Basic of BBO

Practical work about biogeography started their research in 1960s by Alfred Wallance [8] and Charles Darwin [9]. In 2008, Dan Simon [5] proposed a new optimization method based on biogeography called BBO, which are widely applied in Engineering Optimization [10-15].
The mathematical model of biogeography demonstrates the process that species migrate from one island to another. Islands or habitats that are well suited for species obtain high Habitat Suitability Index (HSI).
In BBO methodology, every individual has its own immigration rate Im and emigration rate Em. The immigration rate and emigration rate are both the function of number of species in the habitat. The immigration rate and the emigration rate are calculated as follows:

$$
\begin{align*}
& E m_{k}=E . k / n  \tag{1}\\
& \operatorname{Im}_{k}=I .(1-k / n), \tag{2}
\end{align*}
$$

where k is the number of species in the habitat and n is the largest number of species contained in the habitat; E is the largest emigration rate when there are zero species in the habitat; I is the largest immigration rate when there are the largest number of species which can be contained in the habitat; Suppose all the habitats contain the same largest number of species.
Suppose the probability that a habitat has exactly $k$ species is $\operatorname{Pr}_{k}$, and $\operatorname{Pr}_{k}$ changes from time $t$ to time $t+\Delta t$ as follows [5] :
$\operatorname{Pr}_{k}(t+\Delta t)=\operatorname{Pr}_{k}(t) \cdot\left(1-\operatorname{Im}_{k} \cdot \Delta t-E m_{k} \cdot \Delta t\right)+\operatorname{Pr}_{k-1}(t) \cdot \operatorname{Im}_{k-1} \cdot \Delta t+\operatorname{Pr}_{k+1}(t) \cdot E m_{k+1} \cdot \Delta t$,
where $\mathrm{Im}_{\mathrm{k}}$ is immigration rate when the current habitat has k species; $\mathrm{Im}_{\mathrm{k}-1}$ is immigration rate when the current habitat has $\mathrm{k}-1$ species; $\mathrm{Em}_{\mathrm{k}}$ is emigration rate when the current habitat has k species; $\mathrm{Em}_{\mathrm{k}-1}$ is emigration rate when the current habitat has $\mathrm{k}-1$ species; $\mathrm{Pr}_{\mathrm{k}+1}$ is the probability that the current habitat has exactly $\mathrm{k}+1$ species; $\mathrm{Pr}_{\mathrm{k}-1}$ is the probability that the current habitat has exactly $\mathrm{k}-1$ species.
Suppose $\Delta \mathrm{t}$ is small enough such that we can ignore the probability of more than one emigration and immigration. Calculating the limit of equation (3) as $\Delta \mathrm{t} \rightarrow 0$ obtain equation (4):
$\operatorname{Pr}_{k}=\left\{\begin{array}{lc}-\left(\operatorname{Im}_{k}+E m_{k}\right) \cdot \operatorname{Pr}_{k}+E m_{k+1} \cdot \operatorname{Pr}_{k+1}, & k=0 \\ -\left(\operatorname{Im}_{k}+E m_{k}\right) \cdot \operatorname{Pr}_{k}+\operatorname{Im}_{k-1} \cdot \operatorname{Pr}_{k-1},+E m_{k+1} \cdot \operatorname{Pr}_{k+1}, & k \varepsilon\left[1, S_{m a x}-1\right] \\ -\left(\operatorname{Im}_{k}+E m_{k}\right) \cdot \operatorname{Pr}_{k}+\operatorname{Im}_{k-1} \cdot \operatorname{Pr}_{k-1}, & k=S_{\text {max }}\end{array}\right.$
where $S_{\text {max }}$ is the maximum number of species in a single habitat, the rest are the same as equation (3). If a given solution S with k species has a low probability $\mathrm{Pr}_{\mathrm{k}}$, it is amazing that it is regarded as a solution. The mutation rate $\mathrm{m}_{\mathrm{k}}$ is as follows.
$m_{k}=m_{\text {max }} \cdot\left(1-\operatorname{Pr}_{k}\right) / \operatorname{Pr}_{\text {max }}$
Where $\mathrm{m}_{\max }$ is the maximum number of mutation rate.

## Analysis of Convergence rate

We name a population an optimal population, which contains one optimal individual at least. There may be more than one optimal population. An objective subspace $B_{0}^{*}(\subset \Omega)$ is consists of all the optimal populations.
$\Omega$ is the population space. The aim of BBO is to find a population which is in $B_{0}^{*}(\subset \Omega)$.

Definition 1 The optimal value is defined as $S(X)=\min \left\{\cos t\left(x_{i}\right), i=1,2, \ldots, N P\right\}$. An objective subspace
$B_{0}^{*}=\{\boldsymbol{X} \mid s(\boldsymbol{X})=\min \{\cos t(x), x \in S\}\}$, where $S$ is the solution space.

Definition 2 Given a Markov population series $\left\{\xi_{t}, t=0,1,2, \ldots\right\}\left(\xi_{t} \in \Omega\right)$ and An objective subspace $B_{0}^{*} \subset \Omega$, where $\Omega$ is the population space. We define $u_{t}$ as the probability of $\xi_{t}$ being in the objective subspace $B_{0}^{*}$, i.e.,

$$
\begin{equation*}
u_{t}=\sum_{y \in B_{0}^{*}} P\left(\xi_{t}=y\right), \quad t=0,1,2, \ldots \tag{6}
\end{equation*}
$$

Definition 3 Given a Markov population series $\left\{\xi_{t}, t=0,1,2, \ldots\right\}\left(\xi_{t} \in \Omega\right)$ and an objective subspace $B_{0}^{*} \subset \Omega$, where $\Omega$ is the population space[20]. The Markov population series could converge to objective subspace $B_{0}^{*} \subset \Omega$, if

$$
\begin{equation*}
\lim _{t \rightarrow \infty} u_{t}=1 \tag{7}
\end{equation*}
$$

Definition 4 Given a Markov population series $\left\{\xi_{t}, t=0,1,2, \ldots\right\}\left(\xi_{t} \in \Omega\right)$ and an objective subspace $B_{0}^{*} \subset \Omega$, where $\Omega$ is the population space [7]. We define $p_{t}$ as the convergence rate to $B_{0}^{*} \subset \Omega$ at time t , i.e.,

$$
\begin{equation*}
p_{t}=1-\sum_{y \in B_{0}^{*}} P\left(\xi_{t}=y\right), t=0,1,2, \ldots \tag{8}
\end{equation*}
$$

From the Definition 4, we could get a new relationship as follows:

$$
p_{t}=1-u_{t}, t=0,1,2, \ldots
$$

Suppose there is a random variable $\rho$, which denotes the following events:

$$
\begin{aligned}
& \rho=0: X_{0} \in B_{0}^{*} \\
& \rho=1: X_{1} \in B_{0}^{*} \wedge X_{j} \notin B_{0}^{*}(j=0) \\
& \rho=2: X_{2} \in B_{0}^{*} \wedge X_{j} \notin B_{0}^{*}(\forall j \in\{0,1\}) \\
& \rho=3: X_{3} \in B_{0}^{*} \wedge X_{j} \notin B_{0}^{*}(\forall j \in\{0,1,2\}) \\
& \ldots \\
& \rho=t-1: X_{t-1} \in B_{0}^{*} \wedge X_{j} \notin B_{0}^{*}(\forall j \in\{0,1, \ldots, t-2\}) \\
& \rho=t: X_{t} \in B_{0}^{*} \wedge X_{j} \notin B_{0}^{*}(\forall j \in\{0,1, \ldots, t-1\})
\end{aligned}
$$

The random variable $\rho$ indicates the first hitting time, and $E[\rho]$ denotes the average first hitting time (AFHT) or expected first hitting time (EFHT).
Definition 5 The distribution function of the random variable $\rho$ could be defined as

$$
\begin{equation*}
D_{\rho}(t)=P(\rho \leq t)=P(\rho=t)+P(\rho<t) \tag{9}
\end{equation*}
$$

Because the Markov chain in BBO is an absorbing Markov chain, we can obtain

$$
\begin{align*}
u_{t} & -u_{t-1}=\sum_{y \in B_{0}^{*}} P\left(\boldsymbol{X}_{t}=y\right)-\sum_{y \in B_{0}^{*}} P\left(\boldsymbol{X}_{t-1}=y\right) \\
& =P(\rho=t) \tag{10}
\end{align*}
$$

From above equation, we can get

$$
\begin{equation*}
u_{t}=P(\rho=t)+u_{t-1} \tag{11}
\end{equation*}
$$

By comparing Eq.() and Eq.(), we can obtain

$$
\begin{equation*}
u_{t}=D_{\rho}(t) \tag{12}
\end{equation*}
$$

According to definition 4, we can obtain

$$
\begin{equation*}
D_{\rho}(t)=u_{t}=1-p_{t} \tag{13}
\end{equation*}
$$

The following Lemma 3 is to uncover the relationship [7] between AFHT and convergence rate.
Lemma 1 Given two random variable $\alpha$ and $\beta$, where $\mathrm{E}[\alpha]<\infty$ and $\mathrm{E}[\beta]<\infty$, the distribution functions of $\alpha$ and $\beta$ at time t are $D_{\alpha}(t)$ and $D_{\beta}(t)$, respectively, i.e.,

$$
D_{\alpha}(t)=P(\alpha \leq t)=\sum_{j=0}^{t} P(\alpha=j)
$$

$$
D_{\beta}(t)=P(\beta \leq t)=\sum_{j=0}^{t} P(\beta=j) .
$$

$$
\text { if } \forall t \in\{0,1,2,3 \ldots\}, \quad D_{\alpha}(t) \geq D_{\beta}(t), \text { then }
$$

$$
\begin{equation*}
\mathrm{E}[\alpha] \leq \mathrm{E}[\beta], \tag{14}
\end{equation*}
$$

Where $\mathrm{E}[\alpha]=\sum_{t=0}^{\infty} t . P(\alpha=t)<\infty ; \mathrm{E}[\beta]=\sum_{t=0}^{\infty} t \cdot P(\beta=t)<\infty$.
Proof. According to the definition of expected value of random variable $\alpha$, which is $\mathrm{E}[\alpha]=\sum_{t=0}^{\infty} t \cdot P(\alpha=t)$, we can get

$$
\begin{aligned}
& \mathrm{E}[\alpha]=\sum_{t=0}^{\infty} t \cdot P(\alpha=t) \\
& =0 . P(\alpha=0)+\sum_{t=1}^{\infty} t \cdot P(\alpha=t) \\
& =0 \cdot \sum_{t=0}^{0} P(\alpha=0)+\sum_{t=1}^{\infty} t \cdot(P(\alpha \leq t)-P(\alpha \leq t-1)) \\
& =0 \cdot \sum_{t=0}^{0} P(\alpha=0)+\sum_{t=1}^{\infty} t \cdot\left(\sum_{j=0}^{t} P(\alpha=j)-\sum_{j=0}^{t-1} P(\alpha=j)\right) \\
& =0 . D_{\alpha}(0)+\sum_{t=1}^{\infty} t \cdot\left(D_{\alpha}(t)-D_{\alpha}(t-1)\right) \\
& =\sum_{t=1}^{\infty} t \cdot\left(D_{\alpha}(t)-D_{\alpha}(t-1)\right) \\
& =\sum_{j=1}^{+\infty} \sum_{t=j}^{+\infty}\left(D_{\alpha}(t)-D_{\alpha}(t-1)\right) \\
& =\sum_{j=0}^{+\infty} \lim _{t \rightarrow+\infty}\left(D_{\alpha}(t)-D_{\alpha}(j)\right) \\
& =\sum_{j=0}^{+\infty}\left(1-D_{\alpha}(j)\right) .
\end{aligned}
$$

In the above equations, the second and the third equations is according to the definition of distribution function.
The last equation is due to $\lim _{t \rightarrow+\infty} D_{\alpha}(t)=\lim _{t \rightarrow+\infty} \sum_{j=0}^{t} P(\alpha=j)=1$.
By the same way, $\mathrm{E}[\beta]=\sum_{j=0}^{+\infty}\left(1-D_{\beta}(j)\right)$.
To compare $\mathrm{E}[\alpha]$ and $\mathrm{E}[\beta]$, we subtract $\mathrm{E}[\beta]$ with $\mathrm{E}[\alpha]$.

$$
\begin{aligned}
\mathrm{E}[\alpha]-\mathrm{E}[\beta]= & \sum_{j=0}^{+\infty}\left(1-D_{\alpha}(j)\right)-\sum_{j=0}^{+\infty}\left(1-D_{\beta}(j)\right) \\
& =\sum_{j=0}^{+\infty}\left(D_{\beta}(j)-D_{\alpha}(j)\right) \\
& \leq 0
\end{aligned}
$$

The following Lemma 2 is to get the upper bound and the lower bound of convergence rate, which is applied in
discrete space[16].
Lemma 2 Given an absorbing Markov chain $\left\{\boldsymbol{X}_{t}, t=0,1,2, \ldots\right\}\left(\boldsymbol{X}_{t} \in \Omega\right)$ and an objective subspace $B_{0}^{*} \subset \Omega$, where $\Omega$ is the population space. If there are $\left\{\phi_{t}, t=0,1,2, \ldots\right\}$ and $\left\{\varphi_{t}, t=0,1,2, \ldots\right\}$, which can satisfy the following conditions:
(1) $\prod_{t=0}^{+\infty}\left(1-\phi_{t}\right)=0$
(2) $\phi_{t} \leq \sum_{y \notin B_{0}^{*}} P\left(X_{t+1} \in B_{0}^{*} \mid X_{t}=y\right) \frac{P\left(\boldsymbol{X}_{t}=y\right)}{p_{t}} \leq \varphi_{t}$

Then the absorbing Markov chain can reach the objective subspace $B_{0}^{*}(\subset \Omega)$ with a convergence rate $p_{t}$ which is as following

$$
\begin{equation*}
p_{0} \prod_{j=0}^{t-1}\left(1-\varphi_{j}\right) \leq p_{t} \leq p_{0} \prod_{j=0}^{t-1}\left(1-\phi_{j}\right) \tag{15}
\end{equation*}
$$

Proof. According to the definition of $p_{t}$ we can obtain

$$
\begin{aligned}
p_{t}-p_{t-1} & =\left(1-\mu_{t}\right)-\left(1-\mu_{t-1}\right) \\
& =-\left(\mu_{t}-\mu_{t-1}\right) \\
& =-\left(\sum_{y \in B_{0}^{*}} P\left(X_{t}=y\right)-\sum_{y \in B_{0}^{*}} P\left(X_{t-1}=y\right)\right) \\
& =-P(\rho=t) \\
& =-\sum_{y \notin B_{0}^{*}} P\left(X_{t} \in B_{0}^{*} \mid X_{t-1}=y\right) P\left(X_{t-1}=y\right)
\end{aligned}
$$

The above equation is divided by $p_{t-1}$, then we get

$$
\frac{p_{t}-p_{t-1}}{p_{t-1}}=-\sum_{y \notin B_{0}^{*}} P\left(X_{t} \in B_{0}^{*} \mid X_{t-1}=y\right) \frac{P\left(\boldsymbol{X}_{t-1}=y\right)}{p_{t-1}}
$$

Because of the condition (2), we obtain

$$
\begin{aligned}
& -\varphi_{t-1} \leq \frac{p_{t}-p_{t-1} \leq-\phi_{t-1}}{p_{t-1}} \\
& \Rightarrow-p_{t-1} \varphi_{t-1} \leq p_{t}-p_{t-1} \leq-\phi_{t-1} p_{t-1} \\
& \Rightarrow-p_{t-1} \varphi_{t-1}+p_{t-1} \leq p_{t} \leq-\phi_{t-1} p_{t-1}+p_{t-1} \\
& \Rightarrow-\left(1-\mu_{t-1}\right) \varphi_{t-1}+\left(1-\mu_{t-1}\right) \leq 1-\mu_{t} \leq-\phi_{t-1}\left(1-\mu_{t-1}\right)+\left(1-\mu_{t-1}\right) \\
& \Rightarrow\left(1-\mu_{t-1}\right)\left(1-\varphi_{t-1}\right) \leq 1-\mu_{t} \leq\left(1-\mu_{t-1}\right)\left(1-\phi_{t-1}\right) \\
& \Rightarrow p_{t-1}\left(1-\varphi_{t-1}\right) \leq p_{t} \leq p_{t-1}\left(1-\phi_{t-1}\right)
\end{aligned}
$$

The above inequality is unfolded with recursion, and we get

$$
p_{0} \prod_{j=0}^{t-1}\left(1-\varphi_{j}\right) \leq p_{t} \leq p_{0} \prod_{j=0}^{t-1}\left(1-\phi_{j}\right)
$$

Through the above two lemmas, the relationship between AFHT and convergence rate is obtained. the upper and low bound of the convergence rate is also got. Based on the two lemmas [7], the following theorem gives boundary of AFHT.
Theorem 1 Given an absorbing Markov chain $\left\{\boldsymbol{X}_{t}, t=0,1,2, \ldots\right\}\left(\boldsymbol{X}_{t} \in \Omega\right)$ and an objective subspace $B_{0}^{*} \subset \Omega$, where $\Omega$ is the population space. If there are $\left\{\phi_{t}, t=0,1,2, \ldots\right\}$ and $\left\{\varphi_{t}, t=0,1,2, \ldots\right\}$, which can
satisfy the following conditions:
(1) $\prod_{t=0}^{+\infty}\left(1-\phi_{t}\right)=0$
(2) $\phi_{t} \leq \sum_{y \notin B_{0}^{*}} P\left(X_{t+1} \in B_{0}^{*} \mid X_{t}=y\right) \frac{P\left(\boldsymbol{X}_{t}=y\right)}{p_{t}} \leq \varphi_{t}$

Then the absorbing Markov chain can reach the objective subspace $B_{0}^{*}$ with the average first hitting time AFHT, which is as following

$$
\begin{equation*}
p_{0}\left(\varphi_{0}+\sum_{t=2}^{+\infty} t \varphi_{t-1} \prod_{j=0}^{t-2}\left(1-\varphi_{j}\right)\right) \leq \mathrm{E}[\rho] \leq p_{0}\left(\phi_{0}+\sum_{t=2}^{+\infty} t \phi_{t-1} \prod_{j=0}^{t-2}\left(1-\phi_{j}\right)\right) \tag{16}
\end{equation*}
$$

Proof. According to the lemma 2, we get

$$
p_{t} \leq p_{0} \prod_{j=0}^{t-1}\left(1-\phi_{j}\right)
$$

Due to $p_{t}=1-\mu_{t}$, let $t=0$, we get $p_{0}=1-\mu_{0}$.
As the $\mu_{t}$ is essentially the distribution of random variable $\rho$, which is $D_{\rho}(t)=\mu_{t}$. Then the lower bound can be described using $\mu_{t}$

$$
D_{\rho}(t) \geq\left\{\begin{array}{cc}
1-p_{0} & t=0 \\
1-p_{0} \prod_{j=0}^{t-1}\left(1-\phi_{j}\right) & t \neq 0
\end{array}\right.
$$

Suppose there is a flabby variable $\xi$, which is exactly the lower bound of $D_{\rho}$. The expectation of $\xi$ can be calculated as following.

$$
\begin{aligned}
\mathrm{E}[\xi]= & \sum_{t=0}^{\infty} t \cdot P(\xi=t)=0 .\left(1-p_{0}\right)+1 .\left(p_{0}-p_{0}\left(1-\phi_{0}\right)\right) \\
& +\sum_{t=2}^{+\infty} t \cdot\left(p_{0} \prod_{j=0}^{t-2}\left(1-\phi_{j}\right)-p_{0} \prod_{j=0}^{t-1}\left(1-\phi_{j}\right)\right) \\
= & \phi_{0} p_{0}+\sum_{t=2}^{+\infty} t \cdot \phi_{t-1} p_{0} \prod_{j=0}^{t-2}\left(1-\phi_{j}\right) \\
= & p_{0}\left(\phi_{0}+\sum_{t=2}^{+\infty} t \cdot \phi_{t-1}^{t-2}\left(1-\phi_{j=0}^{t}\right)\right)
\end{aligned}
$$

Due to $D_{\rho}(t) \geq D_{\xi}(t)$ and lemma 1, we get

$$
\mathrm{E}[\rho] \leq \mathrm{E}[\xi]
$$

Then we get the upper bound of AFHT

$$
\mathrm{E}[\rho] \leq p_{0}\left(\phi_{0}+\sum_{t=2}^{+\infty} t \cdot \phi_{t-1} \prod_{j=0}^{t-2}\left(1-\phi_{j}\right)\right)
$$

There is a special case when the initial population does not contain the optimal solution. In such case, $p_{0}=1$. As the initial population in most evolutionary algorithms do not contain the optimal solution, the above equation can be simplified with $p_{0}=1$.
Similarly, we can get the lower bound of AFHT.
In general, there are several mutation operation in evolutionary algorithms, such as MutationConstant, Mutation-1/n and Mutation-random-1.

Definition 6 if every bit in the solution can mutate with constant probability, such mutation is called MutationConstant. (if the candidate solution is binary, then every bit can mutate with probability $p_{c}$; if the candidate solution is coded with d-nary, then every gene can mutate with probability $\frac{p_{c}}{(d-1)}$ )
Definition 7 if every bit in the solution can mutate with probability $1 / \mathrm{n}$, such mutation is called Mutation-1/n. ( if the candidate solution is binary, then every bit can mutate with probability $1 / \mathrm{n}$; if the candidate solution is coded with d-nary, then every gene can mutate with probability $\frac{1}{n(d-1)}$ )
Definition 8 if a gene in the solution is selected randomly, and would mutate, such mutation is called Mutation-random-1. (if the candidate solution is binary, then every bit can mutate with probability $1 / \mathrm{n}$; if the candidate solution is coded with d-nary, then every gene can mutate with probability $\frac{1}{n(d-1)}$ )
If the population size is one, then the population space equals the solution space; If a candidate solution is coded with d-nary, both the population space and the solution space are $d^{n}$.
If the population size is greater than one, then the population space does not equal the solution space; If a candidate solution is coded with d-nary the solution space are $d^{n}$. While the population space is associated with the definition of population, such as unordered population and ordered population.
Definition 9 if a candidate solution is denoted as (.), while a population is denoted as \{.\}. if two populations with the same solution but different arranged order are equivalent, such population is called unordered population.
Definition 10 if a candidate solution is denoted as (.), while a population is denoted as \{.\}. if two populations with the same solution but different arranged order are not equivalent, such population is called ordered population.
To distinguish unordered population with ordered population, the following example is shown. If the candidate solution is coded by binary, then suppose the problem size is three and the population size is 4 . there are two population: $P O P_{1}=\{(000),(111),(101),(001)\}, P O P_{2}=\{(000),(111),(001),(101)\}$. If $P O P_{1}$ and $P O P_{2}$ are unordered populations, then they are equivalent. Otherwise, they are not equivalent.
A population space with problem size n and population size PopSize is provided with $\binom{\mathrm{d}^{n}+\operatorname{PopSize}-1}{$ PopSize }
different populations in the case of unordered population. Otherwise, it has $\mathrm{d}^{n \times \text { PopSize }}$ different populations in the case of ordered population.
Proposition 1 if BBO is adopted to solve a optimization problem with problem size n and population size n and there is no optimal solution in the initial population, then the following lower bound is obtained in the case of migration operation and mutation operation independently
(1) if migration operation and Mutation-Constant operation are adopted, then the lower bound of AFHT is as follows:

$$
\begin{equation*}
\mathrm{E}[\rho]=O\left(\max \left(\frac{1}{n} \times \frac{d-1}{p_{c}\left(1-p_{c}\right)^{n-1}}, \frac{1}{n} \times \frac{1}{\zeta(k)} \times \frac{1}{(1-\zeta(k))^{n-1}}\right)\right), \tag{17}
\end{equation*}
$$

Where $p_{c}$ is a constant; n is dimension number of a candidate solution; d is the code number of every gene of a candidate solution ( d is an integer and $\mathrm{d} \geq 2$ ); $\zeta(k)$ is the normalized immigration rate; where k is the number of species in the habitat.
(2) if migration operation and Mutation- $1 / \mathrm{n}$ operation are adopted, then the lower bound of AFHT is as follows:

$$
\begin{equation*}
\mathrm{E}[\rho]=O\left(\frac{d^{n}}{n^{2}}\right) \tag{18}
\end{equation*}
$$

Where n is dimension number of a candidate solution; d is the code number of every gene of a candidate solution ( d is an integer and $\mathrm{d} \geq 2$ ).

Proof. Because of the independent operation of migration and mutation, suppose $\boldsymbol{X}_{t+1}^{Q}$ is the population after migration operation and $\boldsymbol{X}_{t+1}^{M}$ is the population after migration mutation operation. The following is obtained:

$$
\begin{aligned}
& \sum_{y \notin B_{0}^{*}} P\left(X_{t+1} \in B_{0}^{*} \mid X_{t}=y\right) \frac{P\left(\boldsymbol{X}_{t}=y\right)}{p_{t}} \\
& =\sum_{y \notin B_{0}^{*}} P\left(X_{t+1}^{Q} \in B_{0}^{*} \mid X_{t}=y\right) \frac{P\left(\boldsymbol{X}_{t}=y\right)}{p_{t}} \\
& \quad+\sum_{y \notin B_{0}^{*}} P\left(X_{t+1}^{M} \in B_{0}^{*} \mid X_{t}=y\right) \frac{P\left(\boldsymbol{X}_{t}=y\right)}{p_{t}}
\end{aligned}
$$

According to the features of mutation operation, we have

$$
\sum_{y \notin B_{0}^{*}} P\left(\boldsymbol{X}_{t+1}^{M} \in B_{0}^{*} \mid \boldsymbol{X}_{t}=y\right) \frac{P\left(\boldsymbol{X}_{t}=y\right)}{p_{t}} \leq\left\{\begin{array}{l}
n \times \frac{p_{c}\left(1-p_{c}\right)^{n-1}}{d-1}, \text { Mutation-Constant } \\
n^{2}\left(1-\left(\frac{1}{e}\right)\right)\left(\frac{1}{d^{n}}\right), \text { Mutation-1/n }
\end{array}\right.
$$

As to the migration operation, the analysis is as follows. Compared to the optimal solution, the candidate solution is provided with only one gene which is not the same. In such case, only one gene is needed to immigrate and other genes keep the same. Then, the maximum probability for migration operation to get the optimal solution is as follows

$$
P\left(\boldsymbol{X}_{t+1}^{Q} \in B_{0}^{*} \mid \boldsymbol{X}_{t}=y\right) \leq 1-\left(1-\lambda_{k}^{\prime}\left(1-\lambda_{k}^{\prime}\right)^{n-1}\right)^{n}
$$

where $\lambda_{k}^{\prime}=\lambda_{l}+\left(\lambda_{u}-\lambda_{l}\right)\left(\lambda_{k}-\lambda_{\min }\right) /\left(\lambda_{\max }-\lambda_{\min }\right) \quad$ and $\quad \lambda_{l}=0, \lambda_{u}=1 ; \lambda_{\max }$ and $\lambda_{\min } \quad$ are constants; $\lambda_{k}=I(1-k / n \max ), I=1$ and $n \max$ the largest number of species contained in the habitat; k is the number of species in the habitat. After that, $\lambda_{k}^{\prime}=\left(1-k / n \max -\lambda_{\min }\right) /\left(\lambda_{\max }-\lambda_{\min }\right)$. Because of $n \max$, $\lambda_{\text {max }}$ and $\lambda_{\text {min }}$ are all constants, then $\lambda_{k}^{\prime}$ is denoted by a function $\lambda_{k}^{\prime}=\zeta(k)$ :

$$
\begin{aligned}
& \sum_{y \notin B_{0}^{*}} P\left(X_{t+1}^{Q} \in B_{0}^{*} \mid X_{t}=y\right) \frac{P\left(X_{t}=y\right)}{p_{t}} \\
& \leq \sum_{y \notin B_{0}^{*}}\left(1-\left(1-\zeta(k)(1-\zeta(k))^{n-1}\right)^{n}\right) \frac{P\left(X_{t}=y\right)}{p_{t}} \\
& =\left(1-\left(1-\zeta(k)(1-\zeta(k))^{n-1}\right)^{n}\right) \frac{\sum_{y \notin B_{0}^{*}} P\left(X_{t}=y\right)}{p_{t}} \\
& \left(1-\left(1-\zeta(k)(1-\zeta(k))^{n-1}\right)^{n}\right) \\
& \approx n \times \zeta(k)(1-\zeta(k))^{n-1}
\end{aligned}
$$

Therefore, in the case of Mutation-Constant we have

$$
\begin{aligned}
& \sum_{y \notin B_{0}^{*}} P\left(\boldsymbol{X}_{t+1} \in B_{0}^{*} \mid X_{t}=y\right) \frac{P\left(\boldsymbol{X}_{t}=y\right)}{p_{t}} \\
& \leq n \times \frac{p_{c}\left(1-p_{c}\right)^{n-1}}{d-1}+n \times \zeta(k)(1-\zeta(k))^{n-1}
\end{aligned}
$$

Suppose $\varphi_{t}=n \times \frac{p_{c}\left(1-p_{c}\right)^{n-1}}{d-1}+n \times \zeta(k)(1-\zeta(k))^{n-1}$, according to theorem 1 , the lower bound of AFHT is as follows

$$
\begin{aligned}
\mathrm{E}[\rho] & =\left(n \times \frac{p_{c}\left(1-p_{c}\right)^{n-1}}{d-1}+n \times \zeta(k)(1-\zeta(k))^{n-1}\right)^{-1} \\
& =O\left(\max \left(\frac{1}{n} \times \frac{d-1}{p_{c}\left(1-p_{c}\right)^{n-1}}, \frac{1}{n} \times \frac{1}{\zeta(k)} \times \frac{1}{(1-\zeta(k))^{n-1}}\right)\right) .
\end{aligned}
$$

While in the case of Mutation-Constant we have

$$
\begin{aligned}
& \sum_{y \notin B_{0}^{*}} P\left(\boldsymbol{X}_{t+1} \in B_{0}^{*} \mid \boldsymbol{X}_{t}=y\right) \frac{P\left(\boldsymbol{X}_{t}=y\right)}{p_{t}} \\
& \leq n^{2}\left(1-\left(\frac{1}{e}\right)\right)\left(\frac{1}{d^{n}}\right)+n \times \zeta(k)(1-\zeta(k))^{n-1}
\end{aligned}
$$

Suppose $\varphi_{t}=n^{2}\left(1-\left(\frac{1}{e}\right)\right)\left(\frac{1}{d^{n}}\right)+n \times \zeta(k)(1-\zeta(k))^{n-1}$, according to theorem 1 , the lower bound of AFHT is as follows

$$
\begin{aligned}
\mathrm{E}[\rho] & =\left(n^{2}\left(1-\left(\frac{1}{e}\right)\right)\left(\frac{1}{d^{n}}\right)+n \times \zeta(k)(1-\zeta(k))^{n-1}\right)^{-1} \\
& =O\left(\frac{d^{n}}{n^{2}}\right)
\end{aligned}
$$

## Conclusions

In this paper, we propose some theoretical analysis to lay solid foundation for BBO application. To obtain the average first hitting time (AFHT) of BBO, firstly the relationship between AFHT and convergence rate is discussed, and then the upper bound and the lower bound of convergence rate are put forward. When BBO is adopted to solve an optimization problem with problem size n and population size n , if migration operation and Mutation-Constant or Mutation-1/n operation are adopted, then the lower bound of AFHT is proposed.

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