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Research Article

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A Box-Jenkins Method Based Subset Simulating Model for Daily Ugx-Ngn Exchange Rates

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Abstract A 149-point realization of daily exchange rates of the Uganda shilling (UGX) – Nigerian naira (NGN) from 4^{th} October, 2014 to 1^{st} March, 2015, is analyzed by Box-Jenkins methods. By a new fitting algorithm, it is concluded that the time series follows the subset SARIMA $(1,1,0)x(1,1,0)_7$ model. Daily exchange rates between the two currencies may be simulated or forecasted by the model.

Keywords Ugandan shilling, Nigerian naira, foreign exchange rates, SARIMA models

Introduction

Foreign exchange is a major issue in the discussion of world economy. Any trade relationship between the country Uganda and the country Nigeria is based on the relative value of the Uganda Shilling (UGX) and the Nigerian Naira (NGN). In this write-up the daily exchange rates shall be modelled by Box-Jenkins methods. The particular approach shall be the seasonal autoregressive integrated moving average (SARIMA) approach proposed by Box *et al.* [1].

In recent times, many authors have adopted the SARIMA modelling approach to model real life data. Jianfeng (2013) noticed that SARIMA modelling results in closer forecasts to the real data than dynamic linear modelling in forecasting monthly cases of mumps in Hong Kong. He fitted a SARIMA(2,1,1)x(1,1,1)₁₂ model to the time series [2] . Li et al. (2013) modelled monthly outpatient numbers in China by a SARIMA(0,1,1)x(0,1,1)₁₂ [3]. Kibunja et al. (2014) forecasted monthly precipitation in Mount Kenya region using a SARIMA(1,0,1)x(1,0,0)₁₂ model [4]. Valipour (2015) observed that SARIMA modelling outdid its non-linear counterpart ARIMA in long-term runoff forecasting [5]. Hassan and Mohamed (2015) found that a SARIMA(0,0,5)x(1,0,1)₁₂ was the most adequate in the simulation of monthly rainfall drought in the Gadaref region of Sudan [6]. Gikungu et al. (2015) fitted a SARIMA(0,1,0)x(0,0,1)₄ to quarterly Kenyan inflation rates [7].

The purpose of this write-up is to fit a model to the daily exchange rates of Ugandan shilling (UGX) and Nigerian Naira (NGN). Because of the observed seasonal nature of the series, a SARIMA approach is adopted.

Material and Methods

Data

The data for this work are 149 values of daily UGX / NGN exchange rates of October 4, 2014 through March 1, 2015. They were obtained from the website <u>www.exchangerates.org/UGX-NGN-exchange-rate-history.html</u> accessed on March 2, 2015. These numbers are interpreted as the quantities of NGN per UGX.

Seasonal Autoregressive Integrated Moving Average (SARIMA) Models

The definition of a SARIMA model as proposed by Box *et al.* [1] is as follows:

A stationary time series $\{X_t\}$ is said to follow a multiplicative seasonal autoregressive integrated moving

average model of order p, d, q, P, D, Q, s designated SARIMA(p,d,q)x(P,D,Q)_s if A(L) $\Phi(L^s) \nabla^d \nabla_s^D X_t = B(L)\Theta(L^s)\varepsilon_t$

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(1)

where A(L) is a p-order polynomial in L and is called the autoregressive (AR) operator; B(L) is a q-order polynomial in L and is called the moving average (MA) operator; $\Phi(L)$ is a P-order polynomial in L called the seasonal AR operator; $\Theta(L)$ is a Q-order polynomial in L called the seasonal MA operator. The numbers d and D are the non-seasonal and the seasonal differencing orders respectively. L is the backward shift operator defined by $L^k X_t = X_{t-k}$. the number s is the period of the seasonality of the time series. ∇ and ∇_s are the nonseasonal and the seasonal differencing operators respectively. $\{\varepsilon_t\}$ is a white noise process. SARIMA Modelling

Generally the model (1) is estimated beginning with the determination of the orders: p, d, q, P, D, Q and s. The AR orders p and P are estimated by the non-seasonal and seasonal cut-off lags of the partial autocorrelation function, respectively. Similarly the MA orders q and Q are estimated by the non-seasonal and the seasonal cut-off lags of the autocorrelation function respectively. The seasonal period often suggests itself by the known nature of the series. Otherwise it may be suggestive by the correlogram or an analytical inspection of the series. The differencing orders d and D are such that they sum up to 2 at most.

In this work the subset SARIMA modelling algorithm proposed by Etuk and Ojekudo (2015) shall be used [8]. It is the autoregressive-moving-average-duality-based version of the algorithm of Suhartono [9].

Suhartono's algorithm is as follows:

Fit to $\{X_t\}$ the following SARIMA $(0,0,1)x(0,0,1)_s$ model

(2)

(3)

 $X_t = \epsilon_t + \beta_1 \epsilon_{t-1} + \dots + \beta_s \epsilon_{t-s} + \beta_{s+1} \epsilon_{t-s-1}$ If $\beta_{s+1} = 0$ then the model is said to be additive. Otherwise if $\beta_{s+1} = \beta_1 \beta_s$, then the model is said to be multiplicative. Otherwise it is said to be subset.

Etuk and Ojekudo's algorithm which is the dual version of (2) is as follows:

Fit to $\{X_t\}$ the following SARIMA(1,0,0)x(1,0,0)_s model

 $X_t + \alpha_1 X_{t-1} + \dots + \alpha_s X_{t-s} + \alpha_{s+1} X_{t-s-1}$

If $\alpha_{s+1} = 0$ the model is said to be additive. If not, if $\alpha_{s+1} = \alpha_1 \alpha_s$, the model is said to be multiplicative. Otherwise it is said to be subset. Additivity is ascertained if

$$\widehat{\alpha_{s+1}} < 2SE(\widehat{\alpha_{s+1}})$$

where SE(.) is the standard error of and ^ denotes the estimate of Multiplicativity is ascertained if

 $\mathbf{T} = (\widehat{\alpha_{s+1}} - \widehat{\alpha_s} \ \widehat{\alpha_1} \) / SE(\widehat{\alpha_{s+1}})$

is not statistically significant where T is t-distributed.

Estimation of the model parameters is done via a non-linear optimization process for the mixed ARMA process. The Eviews software which uses the least squares technique is to be used for this work.

Results and Discussion

The time-plot of Figure 1 shows a generally positive trend depicting relative depreciation of the

Naira within the time period of interest. A seven-day differencing yields a series with the time-plot of Figure 2 which depicts a generally horizontal trend and a correlogram of figure 3 showing a seasonal nature of period 7 days. A further non-seasonal differencing yields a series with the plot of Figure 4 and the correlogram of Figure 5. Evident is a stationary nature and a correlation structure suggestive of a SARIMA $(1,1,0)x(1,1,0)_7$. Applying the algorithm of Etuk and Ojekudo (3) becomes naturally suggestive [8].

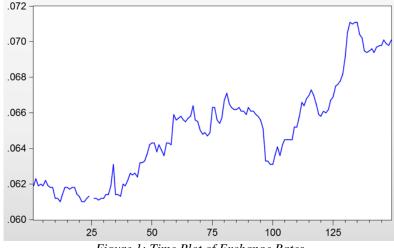
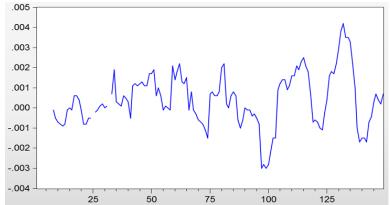
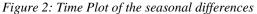
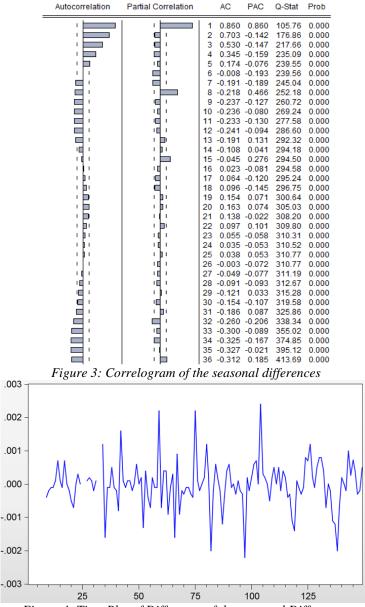


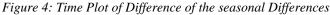
Figure 1: Time Plot of Exchange Rates













 $\begin{array}{l} \mbox{Applying algorithm (3) yields as estimated in Table 1 the SARIMA(1,1,0)x(1,1,0)_7 model \\ X_t = 0.1616X_{t-1} - 0.5983X_{t-7} + 0.1802X_{t-8} + \epsilon_t \\ (\pm 0.0932) \quad (\pm 0.0740) \quad (\pm 0.0920) \end{array}$

Clearly the model is neither additive nor multiplicative. It is subset. From Figure 6, by the Jarque-Bera test, the residuals may be said to follow a normal distribution at 1% level of significance. This implies that the model may be considered as adequate.

Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob
	ı <u>þ</u> ı	1 0.064	0.064	0.5813	0.446
1 🛛 1	1 11	2 0.051	0.047	0.9430	0.624
1 🛛 1	1 1	3 0.053	0.048	1.3498	0.717
i d i	10	4 -0.044	-0.053	1.6266	0.804
ւիւ	լին	5 0.045	0.047	1.9243	0.860
1 1	1 1	6 0.007	0.004	1.9321	0.926
		7 -0.570	-0.578	49.483	0.000
i (fi i	i])i	8 -0.026	0.061	49.586	0.000
ı 🖬 i	1 1	9 -0.097	-0.042	50.986	0.000
1 1	iĝi	10 0.002	0.076	50.987	0.000
i 🏻 i	111	11 0.032	-0.017	51.145	0.000
□ □ '	⊑ !	12 -0.208	-0.217	57.731	0.000
יםי	יםי	13 -0.135	-0.133	60.542	0.000
i pi		14 0.066	-0.353	61.225	0.000
111	ון ו	15 -0.017		61.268	0.000
ים	ון ו	16 0.118	0.071	63.458	0.000
1) 1	יםי	17 0.010	0.114	63.473	0.000
יםי		18 -0.101	-0.163	65.118	0.000
· 🗖 ·	יםי	19 0.186	-0.106	70.690	0.000
i þ	וםי	20 0.133	-0.056	73.558	0.000
ים	[21 0.057	-0.180	74.089	0.000
111	ון ו	22 0.016	0.048	74.132	0.000
יםי	ון ו	23 -0.074	0.059	75.036	0.000
יםי	יםי	24 -0.082	-0.105	76.157	0.000
· 🗖 ·	1 1	25 0.172	-0.002	81.170	0.000
1 1	I]I	26 0.002		81.171	0.000
1 1	ון ו	27 -0.006		81.178	0.000
ı 🛛 i	יםי	28 -0.060	-0.101	81.803	0.000
1) 1	ון ו	29 0.010	0.073	81.822	0.000
1) 1	וםי	30 0.010	-0.066	81.839	0.000
i þ	וםי	31 0.152	0.113	86.013	0.000
יםי	1 1	32 -0.137	-0.003	89.412	0.000
i (Li	וים	33 -0.031	0.115	89.588	0.000
יםי	יםי	34 -0.108	-0.083	91.738	0.000
ı 🛛 i	🗖 '	35 -0.050	-0.262	92.213	0.000
111	([)	36 -0.011	-0.035	92.235	0.000

Figure 5: Correlogram of Difference of the Seasonal Differences

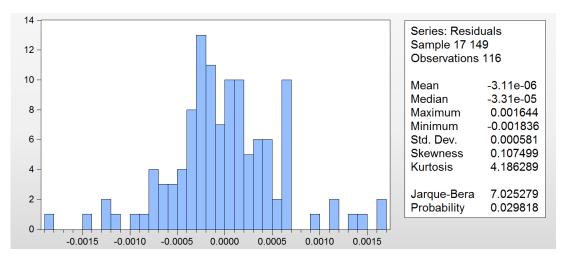


Figure 6: Histogram of The Sarima(1,1,0)X(1,1,0)₁₂ Residuals



 Table 1: Estimation of the Sarima(1,1,0)X(1,1,0)7 Model

 Method: Least Squares

 Date: 10/03/15

 Date: 10/03/15

 Time: 16:09

 Sample (adjusted): 17

 Included observations: 116 after adjustments

 Convergence achieved after 3 iterations

Variable	Coefficient	Std. Error	t-Statistic	Prob.
AR(1)	0.161555	0.093165 1.734076		0.0856
AR(7)	-0.598328	0.074033	0.074033 -8.081868	
AR(8)	0.108211	0.092019	1.175960	0.2421
R-squared	0.373022	Mean dependent var		6.90E-06
Adjusted R-squared	0.361926	S.D. dependent var		0.000733
S.E. of regression	0.000586	Akaike info criterion		-12.02211
Sum squared resid	3.88E-05	Schwarz criterion		-11.95090
Log likelihood	700.2825	Hannan-Quinn criter.		-11.99320
Durbin-Watson stat	1.984381			
Inverted AR Roots	.8340i	.83+.40i	.20+.91i	.2091i
	.18	58+.73i	5873i	93

Conclusion

It may be concluded that daily UGX-NGN exchange rates follow a subset $SARIMA(1,1,0)x(1,1,0)_7$ model. Forecasting and simulation of the series may therefore be based on the proposed model (4).

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